

Kriging of financial term-structures

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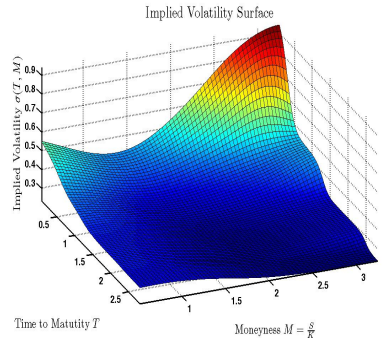
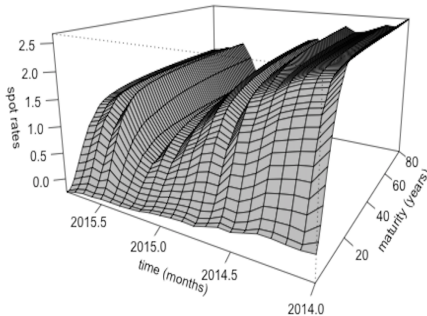
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Motivation

- Financial term-structures describes the evolution of some financial or economic quantities as a function of time horizon.
- **Examples** : interest-rates, bond yields, credit spreads, implied default probabilities, implied volatilities.
- **Applications** : valuation of financial and insurance products, risk management



The term-structure construction problem

Several constraints have to be considered

- **Compatibility with market information** : at a given date t_0 , the curve under construction $T \rightarrow P(t_0, T)$ shall be compatible with observed prices of some reference products.
- **Arbitrage-free construction** : this translates into some specific shape properties such as positivity, monotonicity, convexity or bounds on the curve values
- **Additional conditions can be required** : minimum degree of smoothness, control of local convexity

The term-structure construction problem

1) Compatibility with market information :

- At time t_0 , we observe the market quotes $S_1, \dots, S_n$ of n liquidly traded instruments
- The values of these products depend on the value of the curve at m input locations $X = (\tau_1, \dots, \tau_m)$

The vector of output values $P(t_0, X) := (P(t_0, \tau_1), \dots, P(t_0, \tau_m))^T$ satisfies a linear system of the form

$$A \cdot P(t_0, X) = \mathbf{b},$$

where

- A is a $n \times m$ real-valued matrix
- $\mathbf{b}$ is a n -dimensional column vector

$n < m \implies$ indirect and partial information on the curve values at $\tau_1, \dots, \tau_m$

2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$ is typically a monotonic bounded function

Range of arbitrage-free OIS discount curves

We observe the quoted par rates S_i of an OIS with maturities T_i , $i = 1, \dots, n$

1) Compatibility with market quotes :

The curve $T \rightarrow P(t_0, T)$ of **OIS discount factors** is such that

$$S_i \sum_{k=1}^{p_i-1} \delta_k P(t_0, t_k) + (S_i \delta_{p_i} + 1) P(t_0, T_i) = 1, \quad i = 1, \dots, n$$

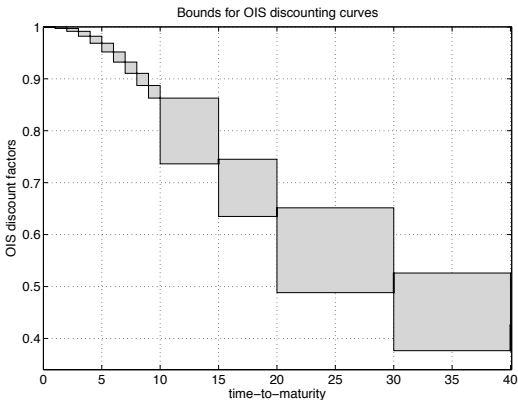
- $t_1 < \dots < t_{p_i} = T_i$: fixed-leg payment dates (annual time grid)
- δ_k : year fraction of period (t_{k-1}, t_k)

2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$ is a decreasing function such that $P(t_0, t_0) = 1$

Range of arbitrage-free OIS discount curves

- $n = 14$ liquidly traded maturities 1, 2, ..., 10, 15, 20, 30, 40 years.
- $m = 40$ points involved in the market-fit linear system
- No-arbitrage bounds on OIS discount factors



Input data : OIS swap rates as of May, 31st 2013.

Source : Cousin and Niang (2014)

Range of arbitrage-free CDS-implied survival functions

We observe at time t_0 the fair spreads S_i of a CDS with maturities T_i , $i = 1, \dots, n$

1) Compatibility with market quotes :

The curve $T \rightarrow P(t_0, T)$ of (risk-neutral) **survival probabilities** is such that

$$S_i \sum_{k=1}^{P_i} \delta_k D(t_0, t_k) P(t_0, t_k) = -(1 - R) \int_{t_0}^{T_i} D(t_0, u) dP(t_0, u), \quad i = 1, \dots, n$$

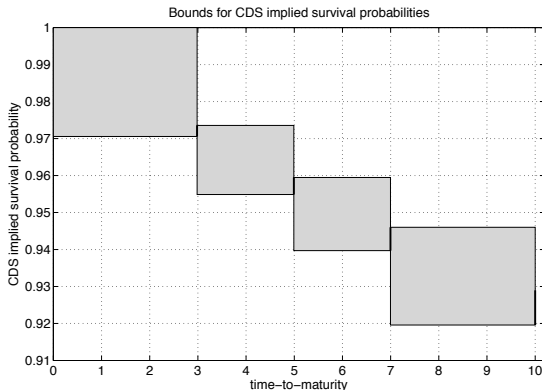
- $t_1 < \dots < t_p = T_i$: trimestrial premium payment dates, δ_k : year fraction of period (t_{k-1}, t_k)
- $D(t_0, T)$ is the discount factor associated with maturity date T
- R : expected recovery rate of the reference entity

2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$ is a decreasing function such that $P(t_0, t_0) = 1$

Range of arbitrage-free CDS-implied survival functions

- $n = 4$ liquidly traded maturities 3, 5, 7, 10 years.
- $m = 40$ points involved in the market-fit linear system
- No-arbitrage bounds on the issuer implied survival distribution function



Input data : CDS spreads of AIG as of December 17, 2007, $R = 40\%$,
 $D(t, T) = \exp(-3\%(T - t))$

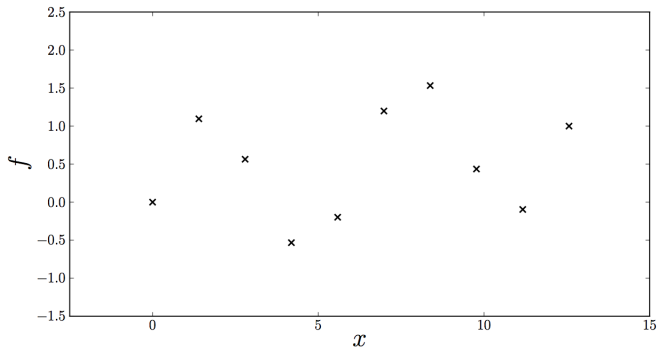
From spline interpolation to kriging

In practice, financial term-structures are constructed using deterministic interpolation techniques.

- Parametric approaches : [Nelson-Siegel](#) or [Svensson](#) models (used by most central banks)
- Non-parametric interpolation methods : shape-preserving spline techniques (lack of interpretability but better ability to fit the data).

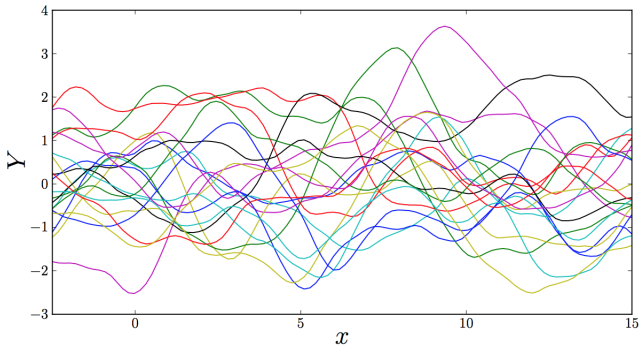
Could we propose an arbitrage-free interpolation method that additionally allows for quantification of uncertainty ?

A function f is only known at a limited number of points $x_1, \dots, x_n$



Classical kriging

The (unknown) function f is assumed to be a sample path of a **Gaussian process** Y



Definition : Gaussian process (GP) or Gaussian random field

A Gaussian process is a collection of random variables, any finite number of which have (consistent) joint Gaussian distributions.

A Gaussian process $(Y(x), x \in \mathbb{R}^d)$ is characterized by its **mean function**

$$\mu : x \in \mathbb{R}^d \longrightarrow \mathbb{E}(Y(x)) \in \mathbb{R}.$$

and its **covariance function**

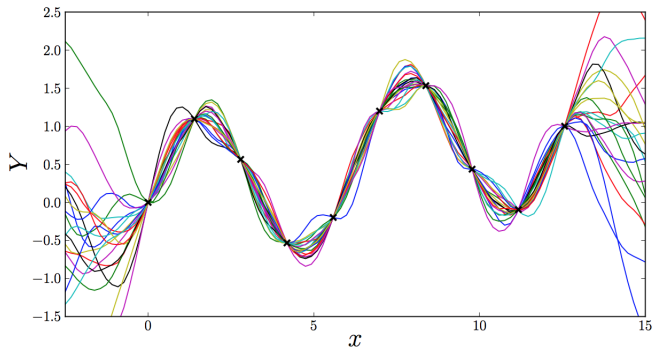
$$K : (x, x') \in \mathbb{R}^d \times \mathbb{R}^d \longrightarrow \text{Cov}(Y(x), Y(x')) \in \mathbb{R}.$$

Table: Some popular covariance functions $K(x, x')$ used in 1D kriging methods.

Name	Expression	Class
Gaussian	$\sigma^2 \exp\left(-\frac{(x-x')^2}{2\theta^2}\right)$	$\mathcal{C}^\infty$
Matérn 5/2	$\sigma^2 \left(1 + \frac{\sqrt{5} x-x' }{\theta} + \frac{5(x-x')^2}{3\theta^2}\right) \exp\left(-\frac{\sqrt{5} x-x' }{\theta}\right)$	$\mathcal{C}^2$
Matérn 3/2	$\sigma^2 \left(1 + \frac{\sqrt{3} x-x' }{\theta}\right) \exp\left(-\frac{\sqrt{3} x-x' }{\theta}\right)$	$\mathcal{C}^1$
Exponential	$\sigma^2 \exp\left(-\frac{ x-x' }{\theta}\right)$	$\mathcal{C}^0$

Classical kriging

The estimation of f relies on the conditional distribution of Y given the observed values $y_i = f(x_i)$ at points x_i , $i = 1, \dots, n$.



- $\mathbf{X} = (x_1, \dots, x_n)^\top \in \mathbb{R}^{n \times d}$: some design points
- $\mathbf{y} = (y_1, \dots, y_n)^\top \in \mathbb{R}^n$: observed values of f at these points
- $\mathbf{Y}(\mathbf{X}) = (Y(x_1), \dots, Y(x_n))^\top$: vector composed of Y at point $\mathbf{X}$

The conditional process is still a Gaussian Process

Let Y be a GP with mean μ and covariance function K . The conditional process $Y \mid \mathbf{Y}(\mathbf{X}) = \mathbf{y}$ is a GP with mean function

$$\eta(x) = \mu(x) + \mathbf{k}(x)^\top \mathbb{K}^{-1}(\mathbf{y} - \boldsymbol{\mu}), \quad x \in \mathbb{R}^d$$

and covariance function $\tilde{K}$ given by

$$\tilde{K}(x, x') = K(x, x') - \mathbf{k}(x)^\top \mathbb{K}^{-1} \mathbf{k}(x'), \quad x, x' \in \mathbb{R}^d$$

where $\boldsymbol{\mu} = \mu(\mathbf{X}) = (\mu(x_1), \dots, \mu(x_n))^\top$, $\mathbb{K}$ is the covariance matrix of $\mathbf{Y}(\mathbf{X})$ and $\mathbf{k}(x) = (K(x, x_1), \dots, K(x, x_n))^\top$

Recall that, in our term-structure construction problem, the (unknown) real function f satisfies some linear equality constraints of the form

$$A \cdot f(X) = \mathbf{b}, \quad (1)$$

where

- A is a given matrix of dimension $n \times m$
- $X = (x_1, \dots, x_m)^\top \in \mathbb{R}^{m \times d}$
- $f(X) = (f(x_1), \dots, f(x_m))^\top \in \mathbb{R}^m$
- $\mathbf{b} \in \mathbb{R}^n$

Extension to linear equality constraints

- $X = (x_1, \dots, x_m)^\top \in \mathbb{R}^{m \times d}$: some design points
- $\mathbf{b} = (b_1, \dots, b_n)^\top \in \mathbb{R}^n$: right-hand side of the linear system
- $Y(X) = (Y(x_1), \dots, Y(x_m))$: vector composed of Y at point X

The conditional process is still a Gaussian Process

Let Y be a GP with mean μ and covariance function K . The conditional process $Y \mid \mathbf{A}Y(X) = \mathbf{b}$ is a GP with mean function

$$\eta(x) = \mu(x) + (\mathbf{A}\mathbf{k}(x))^\top \left(\mathbf{A}\mathbb{K}\mathbf{A}^\top \right)^{-1} (\mathbf{b} - \mathbf{A}\boldsymbol{\mu}), \quad x \in \mathbb{R}^d$$

and covariance function $\tilde{K}$ given by

$$\tilde{K}(x, x') = K(x, x') - (\mathbf{A}\mathbf{k}(x))^\top \left(\mathbf{A}\mathbb{K}\mathbf{A}^\top \right)^{-1} \mathbf{A}\mathbf{k}(x'), \quad x, x' \in \mathbb{R}^d$$

where $\boldsymbol{\mu} = \mu(X) = (\mu(x_1), \dots, \mu(x_m))^\top$, $\mathbb{K}$ is the covariance matrix of $Y(X)$, $\mathbf{k}(x) = (K(x, x_1), \dots, K(x, x_m))^\top$

New formulation of the problem : estimation of an unknown function f given that

$$\begin{cases} A \cdot f(X) = \mathbf{b} \\ f \in \mathcal{M} \end{cases}$$

where $\mathcal{M}$ is the set of (say) non-increasing functions.

Problem : The conditional process is not a Gaussian process anymore.

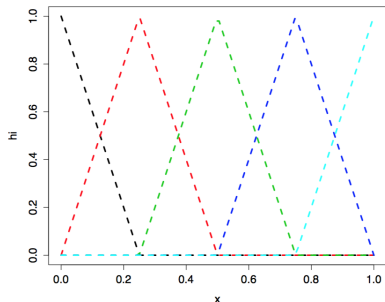
- How to cope with the **infinite-dimensional** monotonicity constraints ?
- Which estimator could we propose for the term-structure ?

Extension to monotonicity constraints (1D case)

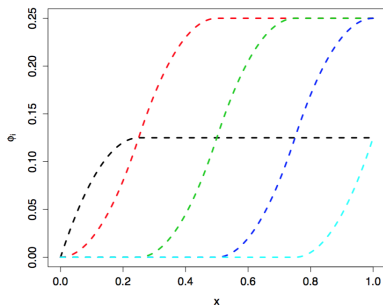
Proposed methodology : On an interval $D = [\underline{x}, \overline{x}]$ of $\mathbb{R}$, we construct a **finite-dimensional approximation** of Y for which the monotonicity constraint is easy to check.

- Regular subdivision $u_0 < \dots < u_N$ of D with a constant mesh δ
- Set of increasing basis functions $(\phi_i)_{i=0, \dots, N}$ defined on this subdivision

$$h_i(x) := \max\left(1 - \frac{|x - u_i|}{\delta}, 0\right)$$



$$\phi_i(x) = \int_{\underline{x}}^x h_i(u) du$$



Extension to monotonicity constraints (1D case)

Proposition (Maatouk and Bay, 2014b)

Let Y be a zero-mean GP with covariance function K and with almost surely differentiable paths.

- The finite-dimensional process Y^N defined on D by

$$Y^N(x) = Y(u_0) + \sum_{j=0}^N Y'(u_j) \phi_j(x)$$

uniformly converges to Y , almost surely.

- Y^N is non-decreasing (resp. non-increasing) on D if and only if $Y'(u_j) \geq 0$ (resp. $Y'(u_j) \leq 0$) for all $j = 0, \dots, N$.
- Let $\xi := (Y(u_0), Y'(u_0), \dots, Y'(u_N))^T$, then $\xi \sim \mathcal{N}(0, \Gamma^N)$ where

$$\Gamma^N = \begin{bmatrix} K(u_0, u_0) & \frac{\partial K}{\partial x'}(u_0, u_j) \\ \frac{\partial K}{\partial x}(u_i, u_0) & \frac{\partial^2 K}{\partial x \partial x'}(u_i, u_j) \end{bmatrix}_{0 \leq i, j \leq N}$$

Extension to monotonicity constraints (1D case)

For a given covariance function K , we assume that the unknown function f is a sample path of the GP

$$Y^N(x) = \eta + \sum_{j=0}^N \xi_j \phi_j(x), \quad x \in D,$$

where $\xi := (\eta, \xi_0, \dots, \xi_N)^\top \sim \mathcal{N}(0, \Gamma^N)$.

Kriging f is equivalent to find the conditional distribution of Y^N given

$$\begin{cases} A \cdot Y^N(X) = \mathbf{b} & \text{linear equality condition} \\ \xi_j \leq 0, j = 0, \dots, N & \text{monotonicity constraint} \end{cases}$$

Extension to monotonicity constraints (1D case)

Or equivalently, to find the distribution of the truncated Gaussian vector $\xi \sim \mathcal{N}(0, \Gamma^N)$ given

$$\begin{cases} A \cdot \Phi \cdot \xi = \mathbf{b} & \text{linear equality condition} \\ \xi_j \leq 0, j = 0, \dots, N & \text{monotonicity constraint} \end{cases}$$

where Φ is a $m \times (N + 2)$ matrix defined as

$$\Phi_{i,j} := \begin{cases} 1 & \text{for } i = 1, \dots, m \text{ and } j = 1, \\ \phi_{j-2}(x_i) & \text{for } i = 1, \dots, m \text{ and } j = 2, \dots, N + 2. \end{cases}$$

Extension to monotonicity constraints (1D case)

Which estimator could we use for f ?

We consider the **mode of the truncated gaussian process** (most probable path) :

$$M_K^N(x \mid A, \mathbf{b}) = \nu + \sum_{j=0}^N \nu_j \phi_j(x),$$

where $\boldsymbol{\nu} = (\nu, \nu_0, \dots, \nu_N)^\top \in \mathbb{R}^{N+2}$ is the solution of the following convex optimization problem :

$$\boldsymbol{\nu} = \arg \min_{\mathbf{c} \in \mathcal{C} \cap \mathcal{I}(A, \mathbf{b})} \left(\frac{1}{2} \mathbf{c}^\top \left(\Gamma^N \right)^{-1} \mathbf{c} \right),$$

with

- $\mathcal{C} = \{ \boldsymbol{\xi} \in \mathbb{R}^{N+2} : \xi_j \leq 0, j = 0, \dots, N \}$
- $\mathcal{I}(A, \mathbf{b}) = \{ \boldsymbol{\xi} \in \mathbb{R}^{N+2} : A \cdot \Phi \cdot \boldsymbol{\xi} = \mathbf{b} \}$

Extension to monotonicity constraints (1D case)

Efficient simulation of the truncated Gaussian vector

1) Simulate a truncated vector ξ given the linear equality constraint :

$$Z \sim \{\xi \mid B \cdot \xi = \mathbf{b}\} \sim \mathcal{N}\left((B\Gamma^N)^\top (B\Gamma^N B^\top)^{-1} \mathbf{b}, \Gamma^N - (B\Gamma^N)^\top (B\Gamma^N B^\top)^{-1} B\Gamma^N\right)$$

where $B = A \cdot \Phi$.

2) Simulate

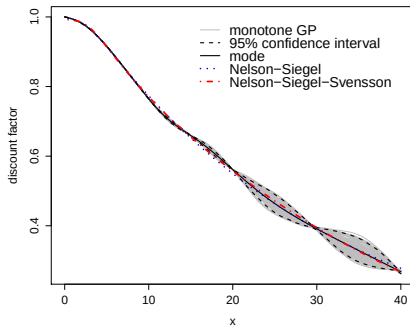
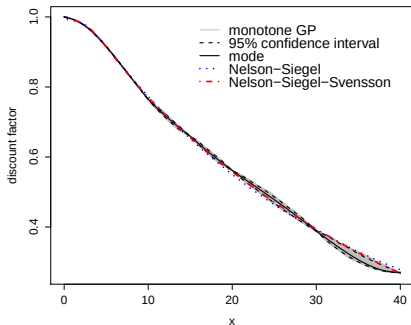
$$\{Z \mid \xi_j \leq 0, j = 0, \dots, N\} \sim \{\xi \mid B \cdot \xi = \mathbf{b} \text{ and } \xi_j \leq 0, j = 0, \dots, N\}$$

by an accelerated rejection sampling method (we use the method proposed in [Maatouk and Bay, 2014a](#))

3) The corresponding sample curves $Y^N(\cdot) = \eta + \sum_{j=0}^N \xi_j \phi_j(\cdot)$ satisfies the constraints on the entire domain D .

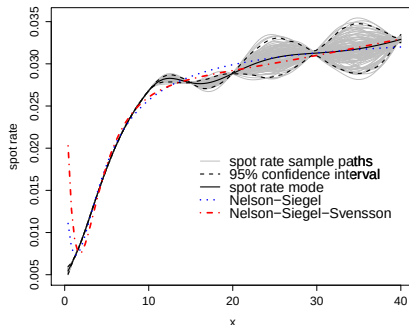
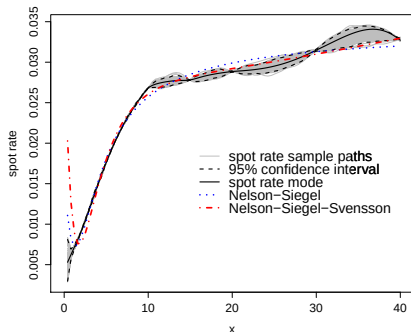
Kriging of OIS discount curves

- We compare two covariance functions : **Gaussian** and **Matérn 5/2**
- Hyper-parameters θ and σ are estimated using cross-validation
- Comparison with **Nelson-Siegel** and **Svensson** curve fitting



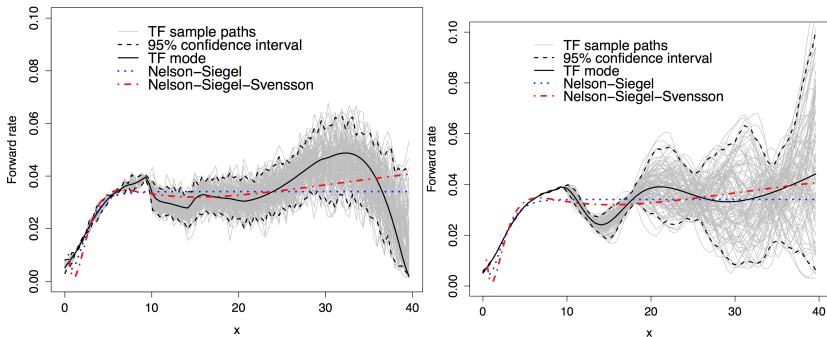
Discount curves. $N = 50, 100$ sample paths. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010.

Corresponding spot rate curves : $-\frac{1}{x} \log P(x)$



Spot rate curves. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010. The black solid line is the most likely spot rate curve $-\frac{1}{x} \log M_K^N(x \mid A, \mathbf{b})$.

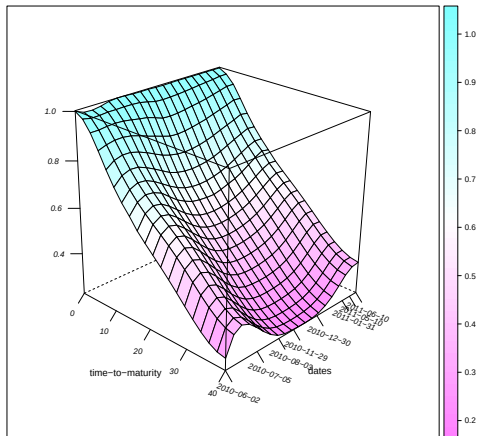
Corresponding forward rate curves : $-\frac{d}{dx} \log P(x)$



Spot rate curves. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010. The black solid line is the most likely forward rate curve $-\frac{d}{dx} \log M_K^N(x | A, b)$.

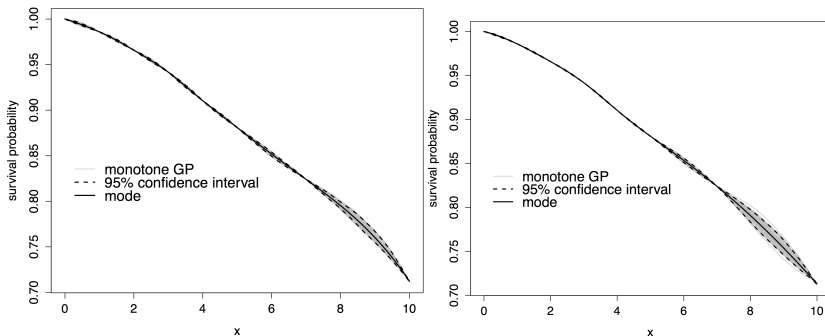
Kriging of OIS discount curves (2D)

The previous approach can be extended in dimension 2.



Discount curves. OIS discount factors as a function of time-to-maturities and quotation dates.

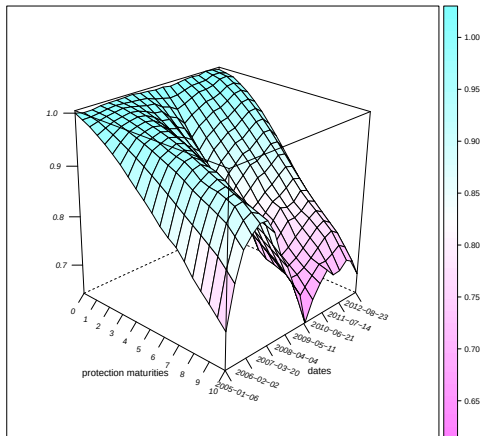
Implied survival function of the Russian sovereign debt



CDS implied survival curves. $N = 50, 100$ sample paths. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. CDS spreads as of 06/01/2005.

Kriging of CDS-implied default distribution (2D)

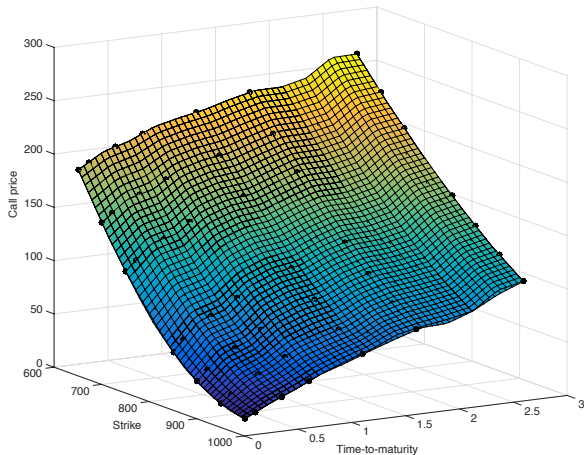
The previous approach can be extended in dimension 2.



Survival curves. CDS implied survival probabilities as a function of time-to-maturities and quotation dates.

- Impact of curve uncertainty on the assessment of related products and their associated hedging strategies
- What if the underlying market quotes are not reliable due to e.g. market illiquidity (data observed with a noise) ?
- Kriging of [arbitrage-free volatility surfaces](#) ?

Kriging of arbitrage-free volatility surface



Thanks for your attention.



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