

Valuation of portfolio loss derivatives in an infectious model

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Empirical studies on contagion mechanisms

- Das and al. (2007) or Azizpour and Giesecke (2008) : **Conditional independence assumption** with **no contagion effect** is rejected by historical default data. The conditional independence assumption is not enough to capture historical default dependency
- Boissay (2006), Jorion and Zhang (2007, 2009) analyze the mechanism of default propagation and provide financial evidence of **chain reactions** or **dominos effects**

Need for a dynamic model with defaults dependencies and contagion

- Eventual underlying macro-economic factors
- Contagion mechanisms
- Chain reactions and evolution over time

Some contagion models in the credit risk field

- Intensities depending on defaults : [Jarrow and Yu \(2001\)](#), [Yu \(2007\)](#)
- Markov chain models : [Schönbucher \(2006\)](#), [Frey and Backhaus \(2007\)](#), [Herbertsson \(2007\)](#), [Laurent, Cousin and Fermanian \(2007\)](#)
- Copula : [Schönbucher and Schubert \(2001\)](#)
- Incomplete information models : [Giesecke \(2004\)](#), [Frey and Runggaldier \(2008\)](#), [Fontana and Runggaldier \(2009\)](#)

In the spirit of Davis and Lo's contagion model

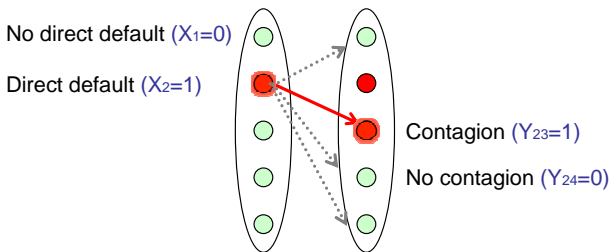
- First models : [Davis and Lo \(2001\)](#)
- Extensions : [Sakata, Hisakado and Mori \(2007\)](#), [Egloff, Leippold and Vanini \(2007\)](#), [Rösch, Winterfeldt \(2008\)](#)
- We propose a multiperiod extension of Davis and Lo's contagion model.

Davis and Lo's contagion model

Modeling of credit contagion for a pool of defaultable entities

- One-period model
- Credit references may default either **directly** or as a consequence of a **contagion effect**

Example : Portfolio with 5 credit references over one period



One-period contagion model

- n : number of credit references,
- X_i : **direct default indicator** of name i (i.e. $X_i = 1$ if i defaults directly, $X_i = 0$ otherwise),
- $Y_{ji} = 1$ if the **contagion link** is activated from name j to name i , $Y_{ji} = 0$ otherwise.

- \mathcal{C}_i : **indirect default indicator** of name i ,
- Z_i : global default indicator (direct or indirect) such that :

$$Z_i = X_i + (1 - X_i)\mathcal{C}_i$$

where :

$$\begin{aligned}\mathcal{C}_i &= \mathbb{1}_{\text{at least one } x_j Y_{ji}=1, j=1, \dots, n} \\ &= 1 - \prod_{j \neq i} (1 - X_j Y_{ji})\end{aligned}$$

$N = \sum_{i=1}^n Z_i$: total number of defaults

Distribution of total number of defaults (Davis and Lo)

$$P[N = k] = C_n^k \sum_{i=1}^k C_k^i p^i (1-p)^{n-i} (1 - (1-q)^i)^{k-i} (1-q)^{i(n-k)}.$$

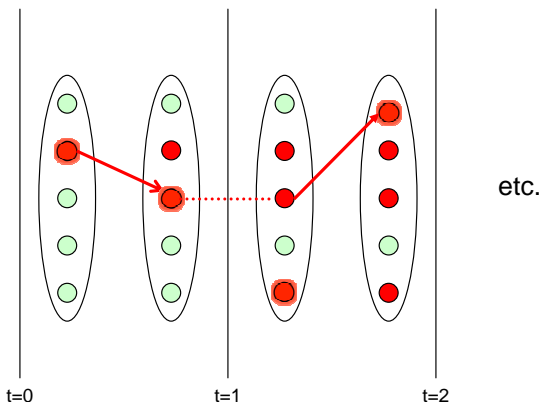
Under the assumptions :

- Direct defaults X_i , $i = 1, \dots, n$: **iid** Bernoulli with parameter p
- Contagion links Y_{ij} , $i, j = 1, \dots, n$: **iid** Bernoulli with parameter q
- **One** contagion link **alone** may trigger an indirect default
- Infected entities cannot contaminate others (**no chain-reaction effect**)

Extension of Davis and Lo's contagion model

Domino Effect

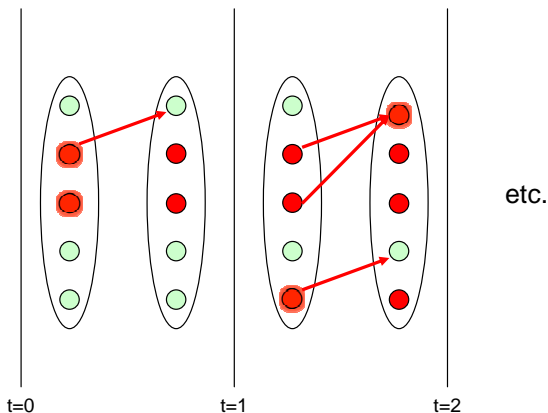
- The model becomes a multiperiod model
- One can choose the set of entities likely to contaminate others
- some iid assumptions are released



Extension of Davis and Lo's contagion model

Contagion incidence on indirect default

- One can change the number of contagion links required to cause a default (here two contaminations required)



Multi-period contagion model : $t = 0, 1, 2, \dots, T$, in period $[t, t + 1]$:

- n : number of credit references,
- X_t^i : **direct default indicator** of entity i ,
- Y_t^{ji} : **contagion links** are Bernoulli random variables such that $Y_t^{ji} = 1$ if entity j may infect entity i ,
- Z_t^i : **global default indicator (direct or indirect)** such that :

$$Z_t^i = Z_{t-1}^i + (1 - Z_{t-1}^i)[X_t^i + (1 - X_t^i)\mathcal{C}_t^i]$$

- $\mathcal{C}_t^i = f\left(\sum_{j \in F_t} Y_t^{ji}\right)$: **indirect default indicator** of name i ,
- F_t is the set of names that are likely to infect other names between t and $t + 1$
- f is a contamination trigger function, for example $f = \mathbb{1}_{x \geq 1}$ (Davis and Lo) or $f = \mathbb{1}_{x \geq 2}$

Extension of Davis and Lo's contagion model

$N_t = \sum_{i=1}^n Z_t^i$: total number of defaults at time t

Main result

$$\begin{aligned} P[N_t = r] &= \sum_{k=0}^r P[N_{t-1} = k] C_{n-k}^{r-k} \sum_{\gamma=0}^{r-k} C_{r-k}^{\gamma} \\ &\quad \cdot \sum_{\alpha=0}^{n-k-\gamma} C_{n-k-\gamma}^{\alpha} \mu_{\gamma+\alpha, t} \sum_{j=0}^{n-r} C_{n-r}^j (-1)^{j+\alpha} \xi_{j+r-k-\gamma, t}(\gamma). \end{aligned}$$

Under the assumptions :

- Direct defaults X_t^i , $i = 1, \dots, n$ are **conditionally independent** Bernoulli r.v. with the same marginal distribution and $\mathbf{X}_t = (X_t^1, \dots, X_t^n)$, $t = 1, \dots, T$ are independent vectors.
- Contagion links Y_t^{ij} , $i, j = 1, \dots, n$ are **conditionally independent** Bernoulli r.v. with the same marginal distribution and $\mathbf{Y}_t = (Y_t^{ij})_{1 \leq i, j \leq n}$, $t = 1, \dots, T$ are independent vectors.
- $(\mathbf{X}_t)_{t=1, \dots, T}$ and $(\mathbf{Y}_t)_{t=1, \dots, T}$ are **independent**.

Similar kind of formulas hold when we have :

finite-exchangeability

- Direct defaults may be **finite-exchangeable** (does not imply conditional independence as infinite exchangeability, De Finetti's Theorem does not apply here).

non stationarity

- Joint law for *Direct defaults* and for *Contagion links* may change over time.

heterogeneity (with higher complexity)

- Direct defaults may be **dependent and heterogeneous**, in a **monoperiodic** framework.
- Direct defaults may be **dependent and heterogeneous**, in a **multi-periodic** framework, but with an exponential complexity (need to consider all possible sets of remaining entities at time t).

Waring's Formula - special case of Schuette-Nesbitt Formula

If B^1, \dots, B^n are n dependent Bernoulli r.v. and $\Gamma \subset \{1, \dots, n\}$ with cardinal m ,

$$\mathbb{P} \left[\sum_{i \in \Gamma} B^i = k \right] = \mathbb{1}_{k \leq m} C_m^k \sum_{j=0}^{m-k} C_{m-k}^j (-1)^j \mu_{j+k}(\Gamma).$$

$$\text{with } \mu_k(\Gamma) = \frac{1}{C_m^k} \sum_{\substack{j_1 < j_2 < \dots < j_k \\ j_1, \dots, j_k \in \Gamma}} \mathbb{P} \left[B^{j_1} = 1 \cap \dots \cap B^{j_k} = 1 \right], \quad k \geq 1,$$

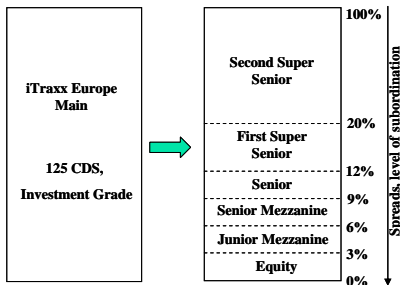
coefficients μ_k may be simplified :

- if independence (without requiring iid) : products
- if exchangeability : the sum vanishes

Here we are looking for :

- Directs defaults : $\sum_{j \in \Gamma} X_t^j$ as a function of some coefficients $\mu_{k,t}(\Gamma)$,
- Contagion links : $\sum_{j \in F_t} Y_t^{\sigma(j)}$ as a function of some coefficients $\lambda_{k,t}$,
- Indirects defaults : $\sum_{j=1 \dots k} \mathcal{C}_t^j$ as a function of some coefficients $\xi_{k,t}$,

Calibration on 5-years iTraxx tranche quotes



- Cash-flows of CDO tranches driven by the **aggregate loss process** (in %)

$$L_t = \frac{1}{n} \sum_{i=1}^n (1 - R_i) Z_t^i$$

where R_i is the **recovery rate** associated with name i .

- if $R_i = R$ for any $i = 1, \dots, n$

$$L_t = \frac{1}{n} (1 - R) \cdot N_t$$

We restrict ourselves to the case where for all t :

- Direct defaults $X_t^i \sim \text{Bernoulli}(\Theta)$ where $\Theta \sim \text{Beta}$, $\mathbb{E}[\Theta] = p$ and $\text{Var}(\Theta) = \sigma^2$, $i = 1, \dots, n$
- Contagion links Y_t^{ij} are iid $Y_t^{ij} \sim \text{Bernoulli}(q)$, $i, j = 1, \dots, n$
- Only one default is required to trigger a default by contagion

Moreover

- $n = 125$, $r = 3\%$ (short-term interest rate)
- Recovery rate $R = 40\%$
- Computation of CDO tranche price only requires marginal loss distributions at several time horizons

Least square calibration procedure : Find $\alpha^* = (p^*, \sigma^*, q^*)$ which minimizes :

$$RMSE(\alpha) = \sqrt{\frac{1}{6} \sum_{i=1}^6 \left(\frac{\tilde{s}_i - s_i(\alpha)}{\tilde{s}_i} \right)^2}.$$

where

	0%-3%	3%-6%	6%-9%	9%-12%	12%-20%	index
Market prices	\tilde{s}_1	\tilde{s}_2	\tilde{s}_3	\tilde{s}_4	\tilde{s}_5	\tilde{s}_0
model prices	$s_1(\alpha)$	$s_2(\alpha)$	$s_3(\alpha)$	$s_4(\alpha)$	$s_5(\alpha)$	$s_0(\alpha)$

Four calibration procedures :

- **Calibration 1** : All available market spreads are included in the fitting
- **Calibration 2** : The equity [0%-3%] tranche spread is excluded
- **Calibration 3** : Both equity [0%-3%] tranche and CDS index spreads are excluded
- **Calibration 4** : All tranche spreads are excluded except equity tranche and CDS index spreads.

Two calibration dates before and during the credit crisis :

- 31 August 2005
- 31 March 2008

Calibration on 5-years iTraxx tranche quotes

31 August 2005

	0%-3%	3%-6%	6%-9%	9%-12%	12%-20%	index
Market quotes	24	81	27	15	9	36
Calibration 1	20	114	7	1	1	29
Calibration 2	-	62	32	18	6	8
Calibration 3	-	55	29	18	7	-
Calibration 4	24	-	-	-	-	36

Annual scaled optimal parameters

	p^*	σ^*	q^*
Calibration 1	0.0016	0.0015	0.0626
Calibration 2	0.0007	0.0133	0.0400
Calibration 3	0.0001	0.0025	0.3044
Calibration 4	0.0014	0.002	0.1090

Calibration on 5-years iTraxx tranche quotes

31 March 2008

	0%-3%	3%-6%	6%-9%	9%-12%	12%-20%	index
Market quotes	40	480	309	215	109	123
Calibration 1	28	607	361	228	95	75
Calibration 2	-	505	330	228	112	68
Calibration 3	-	478	309	215	109	-
Calibration 4	40	-	-	-	-	123

Annual scaled optimal parameters

	p^*	σ^*	q^*
Calibration 1	0.0124	0.0886	0
Calibration 2	0.0056	0.0518	0.0400
Calibration 3	0.0012	0.012	0.2688
Calibration 4	0.0081	0.0516	0.0589

To improve the results we consider :

- One additional external contagious entity
- Calibrated recovery rate

	0%-3%	3%-6%	6%-9%	9%-12%	12%-20%	index	RMSE
31 Jan 2008							
Market spreads	31	317	212	140	74	77	-
Model spreads	32	328	204	142	77	64	7.5
1st Mar 2007							
Market spreads	10	46	13	6	2	23	-
Model spreads	10	37	14	6	2	21	9.2

Tab.: iTraxx Europe main market and model spreads (in bp) and the corresponding root mean square errors. The [0%-3%] spread is quoted in %. All maturities are for five years.

corresponding optimal parameters (on quarterly periods)

	p^*	σ_X^*	q^*	R^*
31 Jan 2008	0.0012	0.0151	0.0007	0.1964
1st Mar 2007	0.0001	0.0026	0.0005	0.1346

Tab.: Optimal parameters $\alpha^* = (p^*, \sigma_X^*, q^*, R^*)$.

We propose a **multi-period extension** of **Davis and Lo's** contagion model that accounts for

- possibly dominos or chain reaction effect
- flexible contagion mechanism (ex : more than one default required to trigger a contamination)
- explicitly model business interdependencies

We provide a **recursive formula** for the **distribution of the number of defaults** at **different time horizons**

- especially when direct defaults and contagion events are **conditionally independent**

The multi-period setting is required to price synthetic CDO tranches

- The contagion parameter has a significant impact on the model ability to fit CDO tranche quotes

I thank you for your attention.

Appendix I - probabilistic tools

Infinite- exchangeability

A_1, A_2, \dots sequence of exchangeable r.v. if for all n and for any permutation σ

$$A_1, \dots, A_n \stackrel{\mathcal{D}}{=} A_{\sigma(1)}, \dots, A_{\sigma(n)},$$

De Finetti's Theorem

A_1, A_2, \dots is a sequence of infinite-exchangeable Bernoulli r.v.

iff there exist a r.v. $\Theta \in [0, 1]$ such that, conditionally to Θ

A_1, A_2, \dots is an iid sequence of Bernoulli r.v. with parameter Θ

- Here, calculations given Θ but difficulties to simplify
- De Finetti's Theorem does not apply for finite-exchangeability
- Need for other tools

If N is a number of fulfilled events B_i , $i \in \Omega$,
A linear combination of $P[N = k]$ will be written :

Schuette-Nesbitt formula

$$\sum_{k \in \Omega} P[N = k] f(k) = \sum_{k \in \Omega} S_k \Delta^k f(0)$$

$$\text{avec } S_k = \sum_{j_1 < \dots < j_k} P[B_{j_1} \cap \dots \cap B_{j_k}]$$

$$\Delta f(k) = f(k+1) - f(k), \text{ difference operator}$$

- events of kind $[N = k]$ given coefficients S_k .
- S_k can be simplified with independence, without requiring i.i.d.
- S_k can be simplified with exchangeability
- events of kind $[N = k]$ as simple as $[N = 0]$ or $[N \geq 1]$

In the particular case where $f(j) = \mathbb{1}_{j=k}$, $j \in \Omega$,

Waring's formula

If X_t^1, \dots, X_t^n are n dependent Bernoulli r.v. and $\Gamma \subset \Omega$ with cardinal m ,

$$P \left[\sum_{i \in \Gamma} X_t^i = k \right] = \mathbb{1}_{k \leq m} C_m^k \sum_{j=0}^{m-k} C_{m-k}^j (-1)^j \mu_{j+k, t}(\Gamma).$$

with

$$\mu_{k, t}(\Gamma) = \frac{1}{C_{\text{card}(\Gamma)}^k} \sum_{\substack{j_1 < j_2 < \dots < j_k \\ j_1, \dots, j_k \in \Gamma}} P \left[X_t^{j_1} = 1 \cap \dots \cap X_t^{j_k} = 1 \right], \quad k \geq 1,$$

$$\mu_{0, t}(\Gamma) = 1 \text{ (even if } \Gamma = \emptyset \text{)}.$$

Interest in life-insurance framework :

- independence assumptions
- but different ages and non identically distributed lifetimes

Interest for Davis and Lo extension :

- one would like $P[N = k]$
- one can change more easily iid assumptions
- is simplified with exchangeability assumptions

Idea from so-called Waring's formula

for non iid Bernoulli r.v. A_1, \dots, A_n , one can get the law of $\sum_j A_j$ as a function of coefficients of kind

$$P[A_1 = 1 \cap \dots \cap A_i = 1].$$

- If independence : these coefficients become products
- If exchangeability : these coefficients does only depend on the number of considered r.v.

Here we are looking for :

- Directs defaults : $\sum_{j \in \Gamma} X_t^j$ as a function of coefficients $\mu_{k,t}(\Gamma)$,
- Contagion links : $\sum_{j \in F_t} Y_t^{\sigma(j)}$ as a function of coefficients $\lambda_{k,t}$,
- Indirects defaults : $\sum_{j=1 \dots k} \mathcal{C}_t^j$ as a function of coefficients $\xi_{k,t}$,

Appendix II - Basic numerical illustration

we consider here that for all t ,

- X_t^i are exchangeables, Bernoulli with hidden parameter Θ_X , $E[\Theta_X] = p = 0.1$, $V[\Theta_X]$ is given
- Y_t^{ij} are exchangeables, Bernoulli with hidden parameter Θ_Y , $E[\Theta_Y] = q = 0.2$, $V[\Theta_Y]$ is given
- hidden parameters are Beta distributed

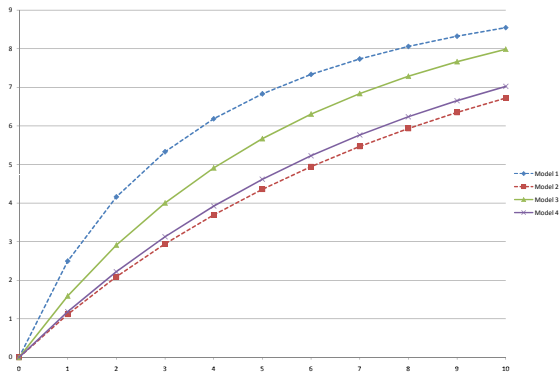
We consider

- 10 entities ($n = 10$),
- 10 temporal units ($T = 10$),
- average direct default probability $p = 0.1$,
- average contagion link probability $q = 0.2$.

We define 4 models with common parameters :

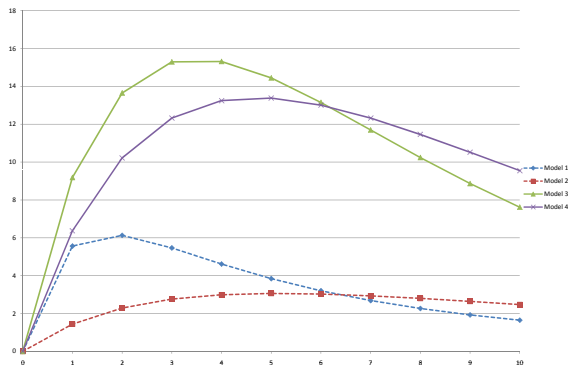
- 1 *model 1* : $\sigma_X = 0, \sigma_Y = 0, f(x) = \mathbb{1}_{x \geq 1}$
(i.i.d. case, one contagion link required).
- 2 *model 2* : $\sigma_X = 0, \sigma_Y = 0, f(x) = \mathbb{1}_{x \geq 2}$
(i.i.d. case, two contagion links required).
- 3 *model 3* : $\sigma_X = 0.2, \sigma_Y = 0.2, f(x) = \mathbb{1}_{x \geq 1}$
(exchangeable case, one contagion link required).
- 4 *model 4* : $\sigma_X = 0.2, \sigma_Y = 0.2, f(x) = \mathbb{1}_{x \geq 2}$
(exchangeable case, two contagion link required).

Appendix II - Basic numerical illustration



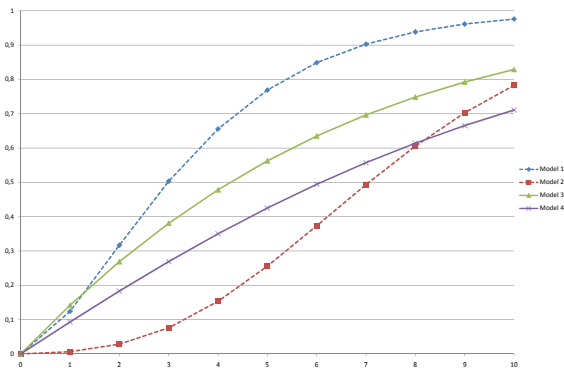
Evolution of $E[N_t]$ as a function of t . i.i.d. case dotted.

Appendix II - Basic numerical illustration



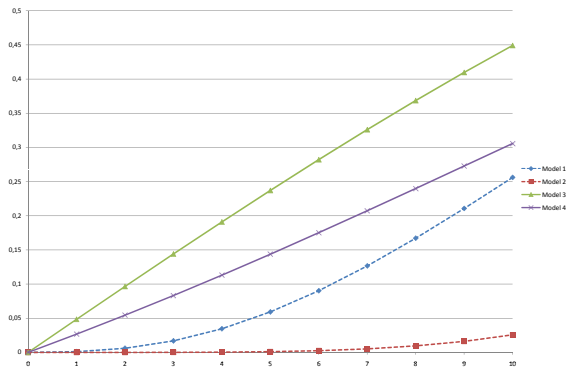
Evolution of $V [N_t]$ as a function of t . i.i.d. case dotted.

Appendix II - Basic numerical illustration



Evolution of $P[N_t \geq 6]$ as a function of t . i.i.d. case dotted.

Appendix II - Basic numerical illustration



Evolution of $P[N_t \geq 10]$ as a function of t . i.i.d. case dotted.

specificity of the model

- try to capture explicit microstructure of contagion
- contagion acts directly on random variables, not on probabilities
- one can say with certainty if default of entity i is due to entity j
- acts in a complete graph

some limits of the model

- default rate depends on the number n of entities
- contagions only within the considered portfolio
- numerical issues for large number n of entities

some perspectives

- recursions to manage numerical issues
- contagions from outside the portfolio
- behavior when time tends to zero and n becomes large
- asymptotic results - larger interconnected component
- recovery effects
- Heterogeneity via a small number of groups