



Some Proposals about Bivariate Risk Measures

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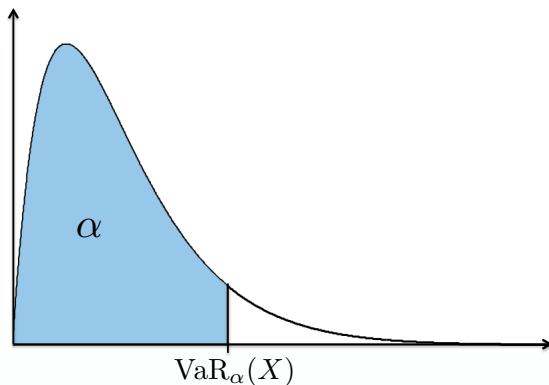
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A. Cousin, E. Di Bernardino, *A multivariate extension of Value-at-Risk and Conditional-Tail-Expectation*

- Under Basel II or Solvency II, each financial institution computes its own regulatory capital in a methodology that does not include risks undertaken by the other institutions even if the latter may be highly interconnected. (“micro-prudential regulation”)
- Risks cannot be diversify away among different institutions. Recent interest for a macro-prudential regulation with an helicopter view on the whole financial system.
- How could we construct a capital rule that reflect both the individual risks and interconnection among these risks in a situation where we cannot benefit from diversification ?

Value-at-Risk paradigm



Given an univariate continuous and strictly monotonic loss distribution function F_X ,

$$\text{VaR}_\alpha(X) = Q_X(\alpha) = F_X^{-1}(\alpha), \quad \forall \alpha \in (0, 1).$$

Shortcoming of VaR measure:

- VaR does not give any information on the severity of loss when larger than the VaR
- VaR is not a coherent risk measure (see Artzner, 1999)

To overcome problems of VaR → Conditional-Tail-Expectation (CTE):

$$CTE_{\alpha}(X) = \mathbb{E}[X | X \geq \text{VaR}_{\alpha}(X)] = \mathbb{E}[X | X \geq Q_X(\alpha)],$$

Dependence and dimensional problems

Riskiness not only of the marginal distributions, but also of the joint distribution:

$$\rho: \mathbf{X} := (X_1, \dots, X_d) \mapsto \begin{pmatrix} \rho^1[\mathbf{X}] \\ \vdots \\ \rho^d[\mathbf{X}] \end{pmatrix} \in \mathbb{R}_+^d,$$

Risk measures essentially based on a “*distributional approach*” (i.e. we have to capture the information coming both from the marginal distributions and from the dependence structure).

Multivariate *Value-at-Risk* as quantile curve (Embrechts & Puccetti, 2006; Nappo & Spizzichino, 2009), i.e., the set

$$\partial L(\alpha) = \{\mathbf{x} \in \mathbb{R}_+^d : F(\mathbf{x}) = \alpha\}$$

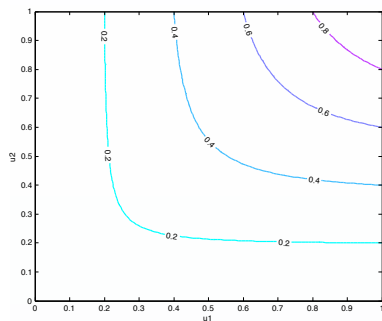
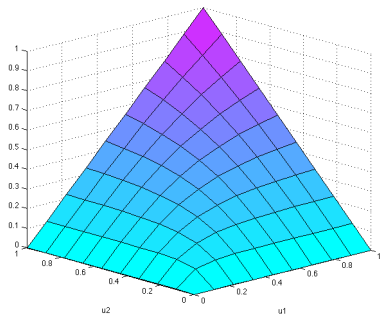


Figure: **left:** cdf of a Clayton copula with parameter 2, **right:** a set of associated quantile curves

A multivariate *Value-at-Risk* and *Conditional-Tail-Expectation*

Definition

Consider a random vector \mathbf{X} with absolutely continuous cdf F . For $\alpha \in (0, 1)$, we define:

$$\text{VaR}_\alpha(\mathbf{X}) = \begin{pmatrix} \mathbb{E}[X_1 | \mathbf{X} \in \partial L(\alpha)] \\ \vdots \\ \mathbb{E}[X_d | \mathbf{X} \in \partial L(\alpha)] \end{pmatrix} = \begin{pmatrix} \mathbb{E}[X_1 | F(\mathbf{X}) = \alpha] \\ \vdots \\ \mathbb{E}[X_d | F(\mathbf{X}) = \alpha] \end{pmatrix},$$

$$\text{CTE}_\alpha(\mathbf{X}) = \begin{pmatrix} \mathbb{E}[X_1 | \mathbf{X} \in L(\alpha)] \\ \vdots \\ \mathbb{E}[X_d | \mathbf{X} \in L(\alpha)] \end{pmatrix} = \begin{pmatrix} \mathbb{E}[X_1 | F(\mathbf{X}) \geq \alpha] \\ \vdots \\ \mathbb{E}[X_d | F(\mathbf{X}) \geq \alpha] \end{pmatrix},$$

where $\partial L(\alpha)$ is the α -level set of F and $L(\alpha)$ is the upper α -level set of F .

Example: Bivariate Archimedean copula case

$$\text{VaR}_{\alpha}^1(X, Y) = \frac{\int_{Q_X(\alpha)}^{\infty} x f_{(U, C(U, V))}(F_X(x), \alpha) dx}{K'(\alpha)},$$

where $f_{(U, C(U, V))}$ is the density of the cdf $F_{(U, C(U, V))}$ given by

$$F_{(U, C(U, V))}(s, t) = t - \frac{\phi(t)}{\phi'(t)} + \frac{\phi(s)}{\phi'(t)}, \quad \text{for } 0 < t < s < 1.$$

and K is the cdf of $C(U, V)$ (Kendall distribution)

Copula	θ	$\text{VaR}_{\alpha, \theta}^1(X, Y)$
Clayton C_{θ}	$(-1, \infty)$	$\frac{\theta}{\theta-1} \frac{\alpha^{\theta}-\alpha}{\alpha^{\theta-1}}$
Counter-monotonic W	-1	$\frac{1+\alpha}{2}$
Independent Π	0	$\frac{\alpha-1}{\ln \alpha}$
Comonotonic M	∞	α

Example: Bivariate Clayton copula case

Remark: for Clayton $\frac{\partial \text{VaR}_{\alpha, \theta}^1}{\partial \alpha} \geq 0$ and $\frac{\partial \text{VaR}_{\alpha, \theta}^1}{\partial \theta} \leq 0$, for $\theta \geq -1$ and $\alpha \in (0, 1)$.

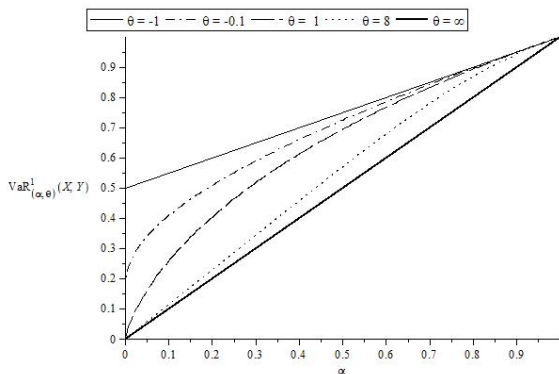


Figure: Behavior of $\text{VaR}_{\alpha, \theta}^1(X, Y) = \text{VaR}_{\alpha, \theta}^2(X, Y)$ with respect to risk level α for different values of dependence parameter θ . The random vector (X, Y) follows a Clayton copula distribution with parameter θ .

Example: Bivariate Clayton copula case

Copula	θ	$\text{CTE}_{\alpha, \theta}^1(X, Y)$
Clayton C_θ	$(-1, \infty)$	$\frac{1}{2} \frac{\theta}{\theta-1} \frac{\theta-1-\alpha^2(1+\theta)+2\alpha^{1+\theta}}{\theta-\alpha(1+\theta)+\alpha^{1+\theta}}$
Counter-monotonic W	-1	$\frac{1}{4} \frac{1-\alpha^2+2 \ln \alpha}{1-\alpha+\ln \alpha}$
Independent Π	0	$\frac{1}{2} \frac{(1-\alpha)^2}{1-\alpha+\alpha \ln \alpha}$
Comonotonic M	∞	$\frac{1+\alpha}{2}$

Table: $\text{CTE}_{\alpha, \theta}^1(X, Y)$ for different copula dependence structures.

Interestingly, one can readily show that $\frac{\partial \text{CTE}_{\alpha, \theta}^1}{\partial \alpha} \geq 0$ and $\frac{\partial \text{CTE}_{\alpha, \theta}^1}{\partial \theta} \leq 0$, for $\theta \geq -1$ and $\alpha \in (0, 1)$.

Example: Bivariate Frank copula case

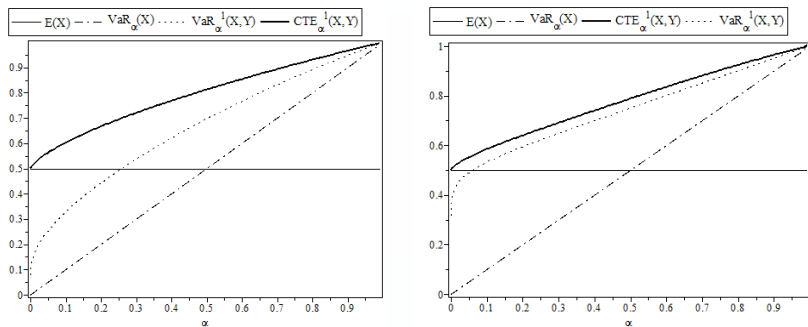


Figure: Frank copula with standard uniform marginals, parameter $\theta = 2$ (left), parameter $\theta = -10$ (right).

	$\text{VaR}_\alpha(\mathbf{X})$	$\text{CTE}_\alpha(\mathbf{X})$
Several (axiomatic) properties	<u>Invariance properties ($\mathbf{c} \in \mathbb{R}_+^d$):</u> <ul style="list-style-type: none"> • $\text{VaR}_\alpha(\mathbf{c}\mathbf{X}) = \mathbf{c} \text{VaR}_\alpha(\mathbf{X})$, • $\text{VaR}_\alpha(\mathbf{c} + \mathbf{X}) = \mathbf{c} + \text{VaR}_\alpha(\mathbf{X})$. <u>Lower bounds:</u> <ul style="list-style-type: none"> • $\text{VaR}_\alpha^i(\mathbf{X}) \geq \text{VaR}_\alpha(X_j), \forall \alpha \in (0, 1)$. <u>Analytical closed-form formulas</u> for $\text{VaR}_\alpha(\mathbf{X})$ and $\text{CTE}_\alpha(\mathbf{X})$.	<u>Invariance properties ($\mathbf{c} \in \mathbb{R}_+^d$):</u> <ul style="list-style-type: none"> • $\text{CTE}_\alpha(\mathbf{c}\mathbf{X}) = \mathbf{c} \text{CTE}_\alpha(\mathbf{X})$, • $\text{CTE}_\alpha(\mathbf{c} + \mathbf{X}) = \mathbf{c} + \text{CTE}_\alpha(\mathbf{X})$. <u>Lower bounds:</u> <ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) \geq \text{VaR}_\alpha(X_j), \forall \alpha \in (0, 1)$. <u>Safety loading:</u> <ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) \geq \mathbb{E}[X_j]$ • $\text{CTE}_0(\mathbf{X}) = \mathbb{E}[\mathbf{X}]$.
Risk level	$\text{VaR}_\alpha^i(\mathbf{X})$ is a non-decreasing function of α .	<ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) \geq \text{VaR}_\alpha^i(\mathbf{X}), \forall \alpha \in (0, 1)$. • $\text{CTE}_\alpha^i(\mathbf{X})$ is a non-decreasing function of α.

- ✓ These two risk measures both satisfy the positive homogeneity and the translation invariance property (Artzner *et al.*, 1999).
- ✓ Comparison results between univariate risk measures and components of multivariate risk measures are provided.
- ✓ Change in risk level α .

	$\text{VaR}_\alpha(\mathbf{X})$	$\text{CTE}_\alpha(\mathbf{X})$
Dependence structure	<p><u>Comonotonic case:</u></p> <ul style="list-style-type: none"> • $\text{VaR}_\alpha^i(\mathbf{X}) = \text{VaR}_\alpha(X_i), \forall \alpha \in (0, 1)$. <p>For a fixed copula C and $X_i \stackrel{d}{=} Y_i$:</p> <ul style="list-style-type: none"> • $\text{VaR}_\alpha^i(\mathbf{X}) = \text{VaR}_\alpha^i(\mathbf{Y}), \forall \alpha \in (0, 1)$. <p>For a fixed copula C and $X_i \leq_{st} Y_i$:</p> <ul style="list-style-type: none"> • $\text{VaR}_\alpha^i(\mathbf{X}) \leq \text{VaR}_\alpha^i(\mathbf{Y}), \forall \alpha \in (0, 1)$. 	<p><u>Comonotonic case:</u></p> <ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) = \text{CTE}_\alpha(X_i), \forall \alpha \in (0, 1)$. <p>For a fixed copula C and $X_i \stackrel{d}{=} Y_i$:</p> <ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) = \text{CTE}_\alpha^i(\mathbf{Y}), \forall \alpha \in (0, 1)$. <p>For a fixed copula C and $X_i \leq_D Y_i$:</p> <ul style="list-style-type: none"> • $\text{CTE}_\alpha^i(\mathbf{X}) \leq \text{CTE}_\alpha^i(\mathbf{Y}), \forall \alpha \in (0, 1)$.

- ✓ Change in marginal distributions and in dependence structure.
- ✓ Results turn to be consistent with existing properties on univariate risk measures.
- ✓ $\theta \leq \theta^* \Rightarrow \text{VaR}_\alpha^1(X^*, Y^*) \leq \text{VaR}_\alpha^1(X, Y)$ (Archimedean copula family).

Perspectives



A. Cousin, E. Di Bernardino, *A multivariate extension of Value-at-Risk and Conditional-Tail-Expectation*, submitted to *Journal of Multivariate Analysis* (2011), <http://hal.archives-ouvertes.fr/hal-00638382/fr/>.

- ✓ Comparisons of our multivariate CTE and VaR with existing multivariate generalizations of these measures, both theoretically and experimentally: applications on financial portfolios; micro-prudential versus macro-prudential approach, ...
- ✓ Extension to the case of discrete distribution functions.

Thank you for your attention